



University of
Salford
MANCHESTER

Subjective well-being and engagement in sport: evidence from England

Forrest, DK and McHale, IG

Title	Subjective well-being and engagement in sport: evidence from England
Authors	Forrest, DK and McHale, IG
Publication title	The economics of sport, health and happiness
Publisher	Edward Elgar
Type	Book Section
USIR URL	This version is available at: http://usir.salford.ac.uk/id/eprint/17951/
Published Date	2011

USIR is a digital collection of the research output of the University of Salford. Where copyright permits, full text material held in the repository is made freely available online and can be read, downloaded and copied for non-commercial private study or research purposes. Please check the manuscript for any further copyright restrictions.

For more information, including our policy and submission procedure, please contact the Repository Team at: library-research@salford.ac.uk.

9. Subjective well-being and engagement in sport: evidence from England

David Forrest and Ian McHale

1. INTRODUCTION

One bishop in the nineteenth-century Anglican Church, whenever asked to say Grace before a public dinner, was wont to pronounce: 'For this food, for our friendship and for all the happiness cricket brings to the World, Thank God.' Implicitly, the Bishop was hypothesizing that subjective well-being is a function of engagement in sport, and it is this possibility that we address in this chapter.

The issue perhaps has more relevance to public policy now than it did back then in the nineteenth century. In Britain, as in many European countries, the contemporary state takes a major part of the responsibility for providing sports facilities: even where it is private clubs, rather than municipalities, that own and operate the centres and complexes, they are often in receipt of grants from lottery funds or quasi-governmental organizations. This steady stream of revenue underpins access to sport in many areas. But, undoubtedly, it is threatened by the current crisis of government debt. Where public expenditure has to be scaled down, sports budgets represent a 'soft' target, partly because it is hard to measure what benefits flow from the subsidies. It therefore appears to us timely to investigate the question of whether people's lives are enhanced by participating in sport and whether it is possible to quantify the benefit.

We take a direct approach. It is argued, for example by Layard (2005), that increasing happiness should be the overriding role of government and, moreover, that this principle provides a practical framework for informing policy decisions because happiness can be measured in a meaningful way by collecting people's happiness scores. Now routinely included in datasets from many social surveys worldwide, happiness scores, often on a 1 to 10 scale, are respondents' answers to a question about how happy their lives are. Patterns of responses to such a question are remarkably similar

across countries (Peiró, 2006) and intuitively appealing (in the sense that things like health and living with a partner are consistently found to be the most important factors contributing to a sense of individual well-being). This has generated increasing confidence among economists that it is legitimate to treat happiness scores as direct measures of the elusive concept of utility and, in examining the association between happiness scores and playing sport, we therefore seek to answer the question of whether government facilitation of participation in sport raises the utility of those who answer the call to play.

Happiness studies is a sub-discipline with a large literature (for a survey from the perspective of economics, see Dolan et al., 2008) but its application to sport, at least through formal econometric analysis, has been very limited. Kavetos and Szymanski (2010) looked at the impact on happiness not of participation but of national team success. Huang and Humphreys (Chapter 8 in this volume) attempt to quantify the causal impact of exercise on self-reported life satisfaction in the USA. Otherwise, there appear to be no prior or contemporaneous studies to report.

2. THE DATA

Taking Part is an annual survey of the use of leisure time by adults (defined as 16 or over) in England. It is commissioned by the UK government's Department for Culture, Media and Sport in conjunction with stakeholder organizations, such as Sport England. We exploit data from the largest-scale edition of this survey, carried out in the 16 months period up to October, 2006. It collected happiness scores from 27 989 adults in England (not Scotland, Wales or Northern Ireland), along with considerable detail on their leisure activities and a range of information on demographic and socio-economic variables. All interviews were conducted face to face and had a median length in excess of half an hour. Fieldwork was by the British Market Research Bureau. The core sample was deemed representative of the adult population (Williams, 2006) but there was also a booster sample of ethnic minority communities to ensure adequate numbers for meaningful analysis of participation in cultural life by race. Data from *Taking Part* are lodged in the UK Data Archive (www.data-archive.ac.uk).

The happiness question posed was similar to others used internationally in general social surveys: 'Taking all things together how happy would you say you are? On a scale of 1–10, 10 = extremely happy, 1 = extremely unhappy.' The phrase 'taking all things together' was intended to encourage respondents to consider not their level of contentment that moment or that day but rather their satisfaction with life more generally. As such,

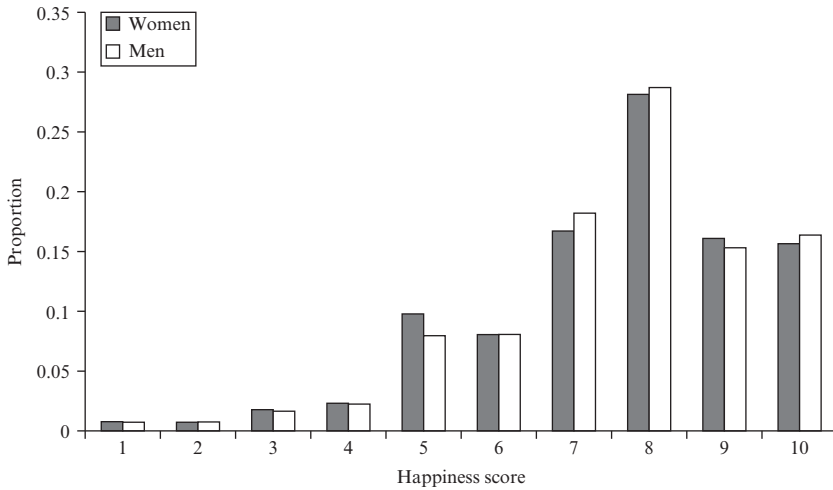


Figure 9.1 Distribution of happiness scores in the sample

answers will be used below not to investigate whether playing sport makes them feel happy on game day (losers of course may not!) but whether there is a more durable effect, in short whether it makes them ‘happier people’.

Fortunately for our purpose, the survey asked not only a happiness question (the distribution of the answers to which are illustrated in Figure 9.1) but also whether subjects had participated in active recreational activities in the four weeks preceding the date of the interview. They were shown a list of 60 activities ranging almost from A to Z (American football to yoga) and were also prompted to mention any others in which they had participated that were not on the list (this uncovered a further eight categories such as frisbee and pilates).

Our focus was to be on sport, so we defined our focus variable *sports player* by reference only to those activities which would conventionally be regarded as sports, as opposed to pastimes or exercise. Of course, a precise definition of sport is hard to formulate (see, for example, Farrell and Shields, 2002) and necessarily involves an element of subjectivity. Nevertheless, we decided that partitioning off ‘sports’ from ‘exercise’ was still appropriate. Sport is a distinctive activity that offers a bundle of characteristics of which exercise is only one: competition (which involves social interaction as well) and skill are other elements which are invariably features of a sport but are not necessarily components of an exercise activity such as jogging. Since they offer different experiences, the relationship between well-being and sport might be misrepresented by examining the

relationship between well-being and an aggregation of sports and exercise activities.

From the list of activities in the survey, many, such as badminton and baseball, cricket and curling, were self-evidently sports. Other pursuits, such as hill trekking and attendance at the gym, we regarded as exercise rather than sport because they lack the element of competition. But a number of cases, including swimming, were ambiguous because they are practised by some participants as keep-fit and by others competitively. Our criterion with these was to treat as 'sport' only those which we judged to be more commonly practised in the context of a contest (the score is kept) than for pure recreation. Accordingly, our final list of 'sports' to be used to construct the indicator variable *sports player* excluded pastimes such as fishing, skiing, swimming and yachting. On our narrow definition, 47.9 per cent of respondents had engaged in sport in the preceding four weeks (a similar order of magnitude to the sports participation rate of 44.5 per cent reported by Farrell and Shields, 2002, from a data-set collected in 1997). Engagement was significantly more common amongst men (54.9 per cent) than among women (42.3 per cent).

3. ARE SPORTS PEOPLE HAPPY PEOPLE?

Mean happiness scores for sports and non-sports players are displayed in Table 9.1. Of course, in calculating means we are implicitly treating the data as cardinal rather than ordinal. But it is not obvious that this is legitimate because the gap in happiness implied by a happiness score of 9 as opposed to 8 might not be equivalent to, say, the gap between 5 and 4. On the other hand, Ferrer-i-Carbonell and Frijters (2004) reviewed evidence

Table 9.1 Sample happiness scores

	Mean happiness score	Standard deviation	Sample size
Whole sample	7.60	1.83	27989
Males	7.63	1.81	12493
Females	7.57	1.85	15496
All sports players	7.73	1.69	13432
All non-sports players	7.48	1.95	14557
Male sports players	7.74	1.69	6871
Male non-sports players	7.50	1.95	5622
Female sports players	7.73	1.70	6561
Female non-sports players	7.46	1.95	8935

consistent with respondents typically answering as if the scale were cardinal. They attributed this to evolution having given people a common understanding of how numbers are used to convey information concerning their feelings, such that they treat the choice of numbers ‘much as they interpret weights in the supermarket’. Fortunately, Ferrer-i-Carbonell and Frijters went on to demonstrate that, in practice, it makes little difference which approach is adopted in studying determinants of happiness, since similar patterns emerged from each when they applied them in turn to major happiness data-sets. Here, we treat the data as cardinal but note that no substantive difference in qualitative findings emerged when we modelled happiness scores employing methods appropriate to ordinal data. We prefer to present ‘cardinal’ results throughout because of their ease of interpretation.

The raw data in Table 9.1 point to sports people being happier on average than non-sports people. The difference of very close to one-quarter of a point is virtually exactly the same whether one considers both genders together or only men or only women. Of incidental interest in Table 9.1 is that the mean and variance of happiness scores of men and women are so close. Here, England may have followed the same path as the USA where the once significant gap in happiness scores (in favour of women) has now entirely disappeared, according to three decades of experience of happiness scores in the *General Social Survey* (Stevenson and Wolfers, 2008).

That the part of the population playing sport self-reports higher happiness than the rest is not in itself very interesting since the difference might be due merely to the composition of the two groups. For example, sports people may also be financially better off and in better health on average than non-sports people. What is interesting is to ask whether there is any difference in happiness score between otherwise similar people according to whether or not they participate. The answer to this question is most conveniently conveyed through ordinary least squares regression.

Table 9.2 presents results from regressing happiness score on the indicator variable *sports player* and a set of control variables selected to reflect gender, age, self-identified ethnicity, labour force status, religious practice, educational qualification, marital status, the presence of children in the household, own evaluation of state of health, income and type of neighbourhood. Most of the variables are self-explanatory (for the groups of indicator variables, the excluded, or reference, categories are *female, white British, employed or not seeking a job, does not practise religion, single, no children present in the household, fair health and other type of neighbourhood*).¹ However, a few require explanation.

Table 9.2 Ordinary least squares regression results

Dependent variable	Happiness score	t
<i>male</i>	-0.011	0.41
<i>age</i>	-0.025	6.42
<i>age squared</i>	0.0003	8.48
<i>Asian</i>	-0.222	5.06
<i>black</i>	-0.070	1.37
<i>Chinese</i>	-0.513	5.37
<i>mixed race</i>	-0.303	6.33
<i>white other</i>		
<i>retired</i>	0.218	5.40
<i>unemployed</i>	-0.460	7.33
<i>student with job</i>	0.297	3.91
<i>student, no job</i>	0.143	2.09
<i>off sick</i>	-0.296	5.28
<i>practises a religion</i>	0.166	5.66
<i>degree</i>	-0.133	5.78
<i>A level</i>	-0.140	3.46
<i>GCSE</i>	-0.130	4.20
<i>other qualification</i>	-0.147	2.71
<i>married</i>	0.674	19.60
<i>cohabiting</i>	0.503	11.30
<i>separated</i>	-0.161	2.55
<i>divorced</i>	-0.053	1.16
<i>widowed</i>	-0.264	5.31
<i>child(ren) aged 8 or below</i>	-0.108	2.96
<i>child(ren) aged 9 to 12</i>	-0.003	0.07
<i>teenager(s)</i>	-0.115	2.52
<i>child(ren) aged 8 or below*male</i>	0.106	1.93
<i>child(ren) aged 9 to 12*male</i>	-0.039	0.57
<i>teenager(s)*male</i>	-0.001	0.02
<i>very good health</i>	1.013	33.05
<i>good health</i>	0.530	18.61
<i>bad health</i>	-0.561	11.48
<i>very bad health</i>	-1.103	5.11
<i>income if reference person</i>	0.003	2.70
<i>not reference person</i>	0.141	2.96
<i>neighbourhood: 'wealthy achievers'</i>	-0.337	1.82
<i>neighbourhood: 'urban prosperity'</i>	-0.449	2.40
<i>neighbourhood: 'comfortably off'</i>	-0.262	1.42
<i>neighbourhood: 'moderate means'</i>	-0.376	2.02
<i>neighbourhood: 'hard pressed'</i>	-0.301	1.62
sports player	0.115	5.07
<i>constant</i>	7.393	35.48

Table 9.2 (continued)

Dependent variable	Happiness score	t
Adjusted R-squared	0.139	
Number of observations	27989	

Note: The specification also included a series of indicator variables to represent missing information on age/health/ethnicity/religious practice/reference person income; there was also an indicator variable for where income was in the top band.

On the magnitude and significance of most of the control predictor variables, there is a strong consensus across studies irrespective of when and where they were conducted. But there have been mixed findings on whether the subjective wellbeing of adults is affected by the presence of children in the household, for example Alesina et al. (2004) found a negative impact. Here we focus on the age rather than the number of children present in the household by including indicator variables to capture the presence or not of one or more children in the young, 9–12 and teenager age groups respectively. Given the possibility that childcare responsibilities press more heavily on women, we also include slope dummy variables (for example, *young child* is multiplied by *male*)² to allow us to show any difference between men and women in how children influence self-reported well-being.

Whether income appears to be an important factor has also varied across studies (for a survey of the relevant literature, see Clark et al., 2008). One of the practical problems in testing is that the same income for an individual could signify poverty (if the level of income is low and the respondent is the breadwinner) or affluence (if the respondent is a secondary earner and can afford to take a part-time or undemanding job). Accordingly, we include income in the model only if the respondent was the 'household reference person' (HRP) (for example, if the person was responsible for paying the rent or mortgage, he or she was the HRP; if partners shared the responsibility, the one with higher income was the HRP). We deem the income of non-HRPs as probably not sufficiently well correlated with family resources to capture living standards and so non-HRPs are represented by setting an appropriate indicator variable to one in their case. It should also be noted that income was collected in the form of bands. We assigned *income*, which refers to the amount before tax received in the preceding 12 months, as the mid-point of the band. This is unlikely significantly to have distorted the results because the bands were narrow, either £2500 or £5000.³

The survey collected the residential postcode of each respondent and,

employing Geographical Information Systems, this could potentially have given access to a rich set of variables relating to neighbourhood characteristics. But, since each UK postcode refers to a very small area (typically 50 or 100 addresses), these were excluded, for reasons of confidentiality, from the data-set released for use by researchers. However, the public data did include an ACORN (A Classification of Residential Neighbourhoods) code supplied by the commercial service CACI which assigns each postcode in the country to one of a series of neighbourhood descriptors (representing the sort of people who live there) on the basis of very micro-level Census statistics concerning the area in which the postcode is located. This was the source of our set of neighbourhood descriptor indicator variables.

Estimation results displayed in Table 9.2 present familiar patterns. Health is by far the most powerful input in the happiness production function; unemployment is associated with quite severely depressed well-being; it is better to have a partner than to be single; most ethnic minorities are less satisfied than white British; well-being is lowest among the middle aged (the turning point in the quadratic relationship between happiness score and age is at age equals 38.05); the religious report themselves happier than others (consistent with Helliwell, 2003); income is a positive factor in determining well-being; and neighbourhood variables play a role (for example, living in an affluent urban area is a negative predictor of well-being; given that the model holds income constant, this could reflect either that satisfaction with income is influenced by levels of neighbours' incomes or that high housing costs erode the amount of discretionary spending). There appears to be no difference in predicted happiness between those who hold different levels of educational qualification (suggesting that benefits from extra qualifications are only achieved through the income variable) but those with no educational qualification at all are happier (income held constant).

There are novel findings on the effect of children. Young children significantly lower happiness scores, though the impact is almost exactly cancelled out by the male slope dummy. Thus we find that the whole cost, in terms of utility, of the burden of looking after young children appears to fall on women. By contrast, both genders seem to suffer equally from a teenager living in the home.

Finally from Table 9.2, the coefficient estimate on our focus variable, *sports player*, is positive and highly significant. Thus, even after controlling for a rich set of variables relating to life circumstances, it remains true that sports players are indeed happier people than non-sports players, albeit that removing composition effects has lowered the difference to a fairly modest level (0.11 happiness points).

4. DOES SPORT MAKE SPORTS PEOPLE HAPPIER PEOPLE?

Ordinary least squares is an effective tool for describing data as it allows us to control for a large number of observed variables. But it cannot allow for the effects of non-observed or unobservable variables and that is why, in the present context, our results so far cannot reasonably be interpreted as showing causation running *from* playing sport *to* well-being: the results are only descriptive. The problem is that the status of *sports player* is not distributed randomly within the sample. Rather, sports players self-selected into that group. If they did so on the basis of non-observed or unobservable personal characteristics that also had a direct impact on happiness score, then ordinary least squares will yield a biased estimate of the *causal* impact of sport on happiness. It will yield an *overestimate* if the relevant personal characteristics that incline an individual to play sport also tend independently to raise his or her ability to achieve happiness. It will yield an *underestimate* if the independent effect of those characteristics is to tend to lower the ability to achieve happiness.

We had no priors concerning the direction of the bias. Let the decision to play sport depend in part on a set of ‘unknown individual characteristics’. These may be ‘favourable’ in the sense of raising an individual’s capacity to achieve happiness or ‘unfavourable’ in the sense of lowering an individual’s ability to achieve happiness. If positive traits, like being tall or having an extrovert personality, dominate in the influence exerted on sports participation by the set of unknown individual characteristics, then persons playing sport will have an atypically high capacity for happiness. In ordinary least squares estimation, there will be omitted variable bias such that the influence of the non-observed or unobservable individual characteristics will be captured in an inflated coefficient estimate on *sports player*. In other words, sports people may be happier because of the sorts of people they are rather than because they play sport. The ordinary least squares coefficient estimate on *sports player* will then be biased upwards.

On the other hand, the decisions of sports people to play sport may have been strongly influenced by negative traits like aggression or over-competitiveness. In this case, those observed to play sport will have a relatively low capacity to achieve happiness and the ordinary least squares estimate will be biased downwards. Sport would be doing more for participants than was evident because sport was attracting disproportionate numbers of people with negative traits. Like the first possibility, this also would not be implausible. For example, aggressive individuals drawn to sport might then avoid behaviour that leads to lower satisfaction with life

if they ‘take out’ their aggression on opponents in a setting where this is socially acceptable.

What is suspected here is a type of selection bias that can be addressed with an appropriate statistical model. The two-step treatment effects model (for a treatment, see Greene, 2008, pp. 889–90) was developed, within the framework of Heckman, for employment in situations such as we face here. The researcher wishes to estimate the effect of treatment (playing sport), represented by a binary variable, on outcome (happiness score). But the decision to undergo treatment is taken by the subject rather than being determined randomly. This raises the possibility of selection bias.

The model proceeds in two stages. The first is a probit to account for the decision whether or not to be treated and the second is the regression of outcome on treatment (and controls). There is selection bias if the errors at stage 1 are correlated with the errors at stage 2. In this event, information extracted at stage 1 is capable of being used at stage 2 to improve explanatory power and, if it is not, there will be omitted variable bias at stage 2. The treatment effects model adds a term, derived from the correlation (if any) between the errors, to capture the influence of selectivity and permit the coefficient estimate on treatment then to be unbiased.

Estimation is enabled to be more precise if it is possible to include exogenous variables at stage 1 which have no direct effect on outcome and can therefore be excluded from stage 2. We employed two such variables, each binary. The first is *access* which is set equal to one for respondents who answered positively to a question on whether sports facilities were available within 20 minutes travelling distance of home (the presence of neighbourhood variables reduces the chance that this will proxy neighbourhood quality). The second is *encouragement* which is set equal to one for respondents who answered positively to a question on whether, as a child, he or she had received parental encouragement to play sport. Unsurprisingly in a small, urbanized country, more than 80 per cent enjoyed *access*. More than 30 per cent had received *encouragement*. We judged that neither variable ‘belonged’ in stage 2 because they would be expected to influence happiness through the *sports player* variable rather than directly.

Table 9.3 displays our results. Some incidental interest attaches to the stage 1 estimates since they illustrate vividly the extent to which sports players in England have a profile that could fairly be described as privileged. Participation is correlated strongly with income, education, belonging to a white ethnic group and residing in an affluent area. Marginal effects here are large. For example, relative to a benchmark subject (defined by continuous variables set to the sample means and indicator variables set to zero), who has a .468 predicted probability of participation, a degree

Table 9.3 Treatment effects model

Dependent variable	Happiness score			
	stage 1 probit	t	stage 2	t
<i>male</i>	0.316	15.28	-0.048	1.41
<i>age</i>	-0.009	2.84	-0.023	5.62
<i>age squared</i>	-0.0001	4.18	0.0003	8.59
<i>Asian</i>	-0.418	12.11	-0.166	3.02
<i>black</i>	-0.343	8.61	-0.023	0.40
<i>Chinese</i>	-0.245	2.10	-0.477	3.10
<i>mixed race</i>	-0.143	2.33	-0.285	3.38
<i>white other</i>	0.009	0.21		
<i>retired</i>	0.134	4.14	0.209	5.11
<i>unemployed</i>	0.100	2.08	-0.471	7.45
<i>student with job</i>	0.108	1.77	0.285	3.72
<i>student, no job</i>	0.155	2.86	0.126	1.81
<i>off sick</i>	-0.243	5.25		
<i>practises a religion</i>	0.014	0.68	0.165	6.47
<i>degree</i>	0.435	15.98	-0.190	3.91
<i>A level</i>	0.341	13.65	-0.185	4.45
<i>GCSE</i>	0.214	8.88	-0.157	4.51
<i>other qualification</i>	0.064	1.49	-0.153	2.81
<i>married</i>	0.006	0.21	0.671	19.42
<i>cohabiting</i>	-0.074	2.15	0.510	11.38
<i>separated</i>	-0.029	0.59	-0.158	2.51
<i>divorced</i>	0.123	3.41	-0.068	1.46
<i>widowed</i>	0.033	0.81	-0.273	1.43
<i>child(ren) aged 8 or below</i>	-0.087	3.13	-0.097	2.60
<i>child(ren) aged 9 to 12</i>	0.072	2.20	-0.014	0.33
<i>teenager(s)</i>	-0.036	1.04	-0.109	2.38
<i>child(ren) aged 8 or below*male</i>	0.026	0.62	0.100	1.81
<i>child(ren) aged 9 to 12*male</i>	0.102	1.90	-0.051	0.73
<i>teenager(s)*male</i>	0.039	0.71	-0.008	0.12
<i>very good health</i>	0.358	15.02	0.969	24.09
<i>good health</i>	0.217	9.74	0.503	15.53
<i>bad health</i>	-0.217	5.09	-0.541	10.75
<i>very bad health</i>	-0.350	3.79	-1.001	10.55
<i>income if reference person</i>	0.004	4.66	0.002	2.05
<i>not reference person</i>	0.078	3.52	0.131	4.49
<i>neighbourhood: 'wealthy achievers'</i>	0.350	2.37	-0.379	2.02

Table 9.3 (continued)

Dependent variable	Happiness score			
	stage 1 probit	t	stage 2	t
<i>neighbourhood:</i> <i>'urban prosperity'</i>	0.238	1.57	-0.477	2.53
<i>neighbourhood:</i> <i>'comfortably off'</i>	0.249	1.69	-0.292	1.57
<i>neighbourhood:</i> 'moderate <i>means'</i>	0.148	1.00	-0.393	2.10
<i>neighbourhood:</i> 'hard <i>pressed'</i>	0.101	0.68	-0.313	1.68
sports player			0.467	2.26
<i>access</i>	0.347	12.60		
<i>encouragement</i>	0.110	6.46		
lambda			-0.215	1.72
constant	-0.447	2.67	-7.212	30.78
Pseudo R-squared	0.155			
Number of observations	27989			

Note: The specification also included a series of indicator variables to represent missing information on age/health/ethnicity/religious practice/reference person income; there was also an indicator variable for where income was in the top band.

is estimated to raise this by more than 17 percentage points, whereas membership of the Asian community lowers it by very close to 16 percentage points. Other features of the table include: that males are more likely than females to play sport; that, for a given age and income of course, the unemployed and the retired and students, provided they do not also have a job, are more likely to take part in sport (presumably they are time rich relative to the benchmark); that young children deter parents' participation but the effect is opposite for the 9–12 age group (we speculate that this is the age when parents and children can go to the sports club together); and that cohabitation lowers participation but marriage does not (our interpretation is that most cohabiters are at a stage in their relationship when they still need to be together most of the time). Encouragingly, our two 'additional' variables each appear to exert a strong impact on probability of participation with marginal effects of more than 13 and more than 4 percentage points for *access* and *encouragement* respectively.

In the stage 2 results, the estimated coefficient on the selectivity correction term is significant, though weakly so ($p = .086$), and negative. The sign

Table 9.4 Reduced-form model

Dependent variable	Happiness score	t
<i>male</i>	-0.005	0.20
<i>age</i>	-0.025	6.50
<i>age squared</i>	0.0003	8.40
<i>Asian</i>	-0.228	5.20
<i>black</i>	-0.073	1.42
<i>Chinese</i>	-0.503	3.31
<i>mixed race</i>	-0.305	3.66
<i>white other</i>	-0.206	2.76
<i>retired</i>	0.221	5.48
<i>unemployed</i>	-0.454	7.24
<i>student with job</i>	0.294	3.87
<i>student, no job</i>	0.149	2.18
<i>off sick</i>		
<i>practises a religion</i>	0.163	6.45
<i>degree</i>	-0.128	3.67
<i>A level</i>	-0.137	4.26
<i>GCSE</i>	-0.129	4.19
<i>other qualification</i>	-0.151	2.78
<i>married</i>	0.671	19.50
<i>cohabiting</i>	0.496	11.15
<i>separated</i>	-0.163	2.60
<i>divorced</i>	-0.050	1.08
<i>widowed</i>	-0.262	5.62
<i>child(ren) aged 8 or below</i>	-0.011	3.04
<i>child(ren) aged 9 to 12</i>	-0.0009	0.02
<i>teenager(s)</i>	-0.113	2.47
<i>child(ren) aged 8 or below*male</i>	0.108	1.97
<i>child(ren) aged 9 to 12*male</i>	-0.035	0.51
<i>teenager(s)*male</i>	-0.003	0.04
<i>very good health</i>	1.024	33.52
<i>good health</i>	0.534	18.80
<i>bad health</i>	-0.562	11.49
<i>very bad health</i>	-1.022	10.93
<i>income if reference person</i>	0.003	2.81
<i>not reference person</i>	0.143	5.04
<i>neighbourhood: 'wealthy achievers'</i>	-0.132	1.79
<i>neighbourhood: 'urban prosperity'</i>	-0.447	2.39
<i>neighbourhood: 'comfortably off'</i>	-0.261	1.41
<i>neighbourhood: 'moderate means'</i>	-0.378	2.03
<i>neighbourhood: 'hard pressed'</i>	-0.303	1.63
access	0.076	2.29

Table 9.4 (continued)

<i>encouragement</i>	0.114	2.57
constant	7.354	35.10
Adjusted R-squared	0.141	
Number of observations	27989	

Note: The specification also included a series of indicator variables to represent missing information on age/health/ethnicity/religious practice/reference person income; there was also an indicator variable for where income was in the top band.

implies negative correlation between stage 1 and stage 2 errors and so it is suggestive of ordinary least squares giving downwardly biased estimates of the impact of playing sport. It underestimates the impact of sport because it fails to take into account that a disproportionate number of those who are sports players are endowed with unfavourable characteristics from the perspective of wanting to be happy. The point estimate on *sports player* is 0.467. This represents the *causal* impact of being a sports player on happiness score. It is a large payoff. Though not by any means as important as marriage, the estimated average effect of constraining a current player not to take part in sport is almost exactly the same as that from moving a currently employed subject into unemployment. The framework for policy debate in the near future is likely to be on closing sports facilities as a contribution to reducing the public sector deficit. Our results illustrate that participation in sport makes a significant contribution to the well-being of many people and that if closures cause them to give up or play less often, then this would imply a high social cost.

Finally, and similar to Huang and Humphreys (Chapter 8 in this volume), we estimate a reduced form model to check directly the importance of sports facilities. The underlying model is that *happiness score* depends on the value of *sports player* (which is endogenous) and the values of control variables. In turn, *sports player* is a function of the same controls and *access* and *encouragement*. The reduced form equation then has *happiness score* depend on all the exogenous variables including *access* and *encouragement*.

The point estimate on *access* in Table 9.4 suggests that maintaining the availability of a sports facility to a randomly selected individual raises the expected happiness score by 0.076 points. This could be regarded as a very strong effect indeed. For example, if a particular sports centre were the only facility serving the needs of a local population of 1000, the loss in aggregate happiness points from closing it would be roughly equivalent to that from 155 'household reference persons' across the country becoming

unemployed (this takes into account both the direct impact of unemployment on happiness score and the effect from receiving the sample mean income associated with unemployment rather than the sample mean income associated with employment). As Huang and Humphreys (Chapter 8 in this volume) imply, the loss could be even greater if there is also an effect (on well-being) through health variables (which will happen if health suffers when participation in sport is reduced or terminated). The policy conclusion of our chapter is that sports infrastructure should not lightly be discarded by government.

NOTES

1. The labour force status categories were allocated from answers to questions about whether the interviewee was employed, retired or a student. There was no explicit category called 'unemployed' but we were able to construct our own *unemployed* variable, set equal to one for those respondents who had not answered 'employed' to the main economic status question (they had not done paid work in the preceding seven days) but answered 'yes' to a subsequent question about whether they had been looking for work or a place on a training scheme in the preceding four weeks. This left the reference category to include the employed and others, such as homemakers, who were neither students nor retired but who were not seeking a job.
2. *Male* of course refers to the gender of the respondent, not that of the child.
3. The top band, over £50 000, was unbounded. We gave the respondents in this small category (3 per cent of the sample) the value of £55 000 but, since this was arbitrary, we also included an indicator variable for this group).

REFERENCES

- Alesina, A., Di Tella, R. and MacCulloch, R. (2004), 'Inequality and happiness: are Europeans and Americans different?', *Journal of Public Economics*, **88**, 209–42.
- Clark, A.E., Frijters, P. and Shields, M.A. (2008), 'Relative income, happiness and utility: an explanation for the Easterlin Paradox and other puzzles', *Journal of Economic Literature*, **46**, 95–144.
- Dolan, P., Peasgood, T. and White, M. (2008), 'Do we really know what makes us happy? A review of the economic literature on the factors associated with subjective well-being', *Journal of Economic Psychology*, **29**, 94–122.
- Farrell, L. and Shields, M.A. (2002), 'Investigating the economic and demographic determinants of sporting participation in England', *Journal of the Royal Statistical Society, Series A*, **165**, 335–48.
- Ferrer-i-Carbonell, A. and Frijters, P. (2004), 'How important is methodology for the estimates of the determinants of happiness?', *Economic Journal*, **114**, 641–59.
- Greene, W.H. (2008), *Econometric Analysis*, 6th edn, Upper Saddle River, NJ: Pearson Education.

- Helliwell, J.F. (2003), 'How's life? Combining individual and national variables to explain subjective well-being', *Economic Modelling*, **20**, 331–60.
- Kavetskos, G. and Szymanski, S. (2010), 'The impact of major sporting events on happiness', *Journal of Economic Psychology*, **31** (2), 158–71.
- Layard, R. (2005), *Happiness: Lessons from a New Science*, London: Penguin Books.
- Peiró, A. (2006), 'Happiness, satisfaction and socio-economic conditions: some international evidence', *Journal of Socio-Economics*, **35**, 348–65.
- Stevenson, B. and Wolfers, J. (2008), 'Happiness inequality in the United States', *Journal of Legal Studies*, **37**, 533–79.
- Williams, J. (2006), *Taking Part: The National Survey of Culture, Leisure and Sport. Final Technical Report*, London: British Market Research Bureau Ltd.